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Takayuki Mizuno
University of Tsukuba
Kazumasa Takei
The Norinchukin Bank
Takaaki Ohnishi
Canon Institute for Global Studies
Tsutomu Watanabe
The University of Tokyo

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Temporal and Cross Correlations in Business News

Takayuki MIZUNO^{1,2}, Kazumasa TAKEI³, Takaaki OHNISHI^{2,4},
Tsutomu WATANABE^{2,4}

¹*Division of Information Engineering, Faculty of Engineering, Information and Systems, University of Tsukuba, Ibaraki 305-8573, Japan*

²*Canon Institute for Global Studies, Tokyo 100-6511, Japan*

³*The Norinchukin Bank, Tokyo 100-8420, Japan*

⁴*Graduate School of Economics, University of Tokyo, Tokyo 113-0033, Japan*

We empirically investigate temporal and cross correlations in the frequency of news reports on companies, using a unique dataset with more than 100 million news articles reported in English by around 500 press agencies worldwide for the period 2003-2009. We find, first, that the frequency of news reports on a company does not follow a Poisson process; instead, it exhibits long memory with a positive autocorrelation for more than one year. Second, we find that there exist significant correlations in the frequency of news across companies. On a daily or longer time scale, the frequency of news is governed by external dynamics, while it is governed by internal dynamics on a time scale of minutes. These two findings indicate that the frequency of news on companies has similar statistical properties as trading volumes or price volatility in stock markets, suggesting that the flow of information through company news plays an important role in price dynamics in stock markets.

§1. Introduction

News is the communication of information about important events. In macroeconomics, quantitative finance, and econophysics, the frequency of news has been studied in order to measure the impact of news on prices and trading volumes in stock markets.^{1),2)} It has been shown by some financial economists that there is only a weak relationship between the daily number of news reports, the volume, and the price return in stock markets.³⁾ On the other hand, in the area of econophysics, it has been observed that market volatility and volume increases immediately after particular news have been reported.⁴⁾⁻⁶⁾ Another approach has been to examine the influence of exogenous shocks, including news reports, on pricing in financial markets using numerical models.⁷⁾ Yet another strand of research has attempted to detect patterns in the flow of information. Refs. 8) and 9) for instance suggest that the frequency of use of specific words in blogs on the internet does not follow a Poisson process, while Ref. 10) shows that using latent Dirichlet allocation, news articles appearing in the New York Times can be classified into several topics.

The aim of this paper is to empirically identify certain statistical properties of the frequency of news, with a special focus on the temporal correlation of the frequency of news on a specific company as well as the cross correlation of news across companies. For this purpose, we employ a dataset of news articles reported by around 500 press agencies worldwide. The dataset – “Reuters NewsScope Archive” – was obtained from Thomson Reuters Corporation. The rest of the paper is organized as follows. Section 2 describes our dataset and shows that there are periodicities in

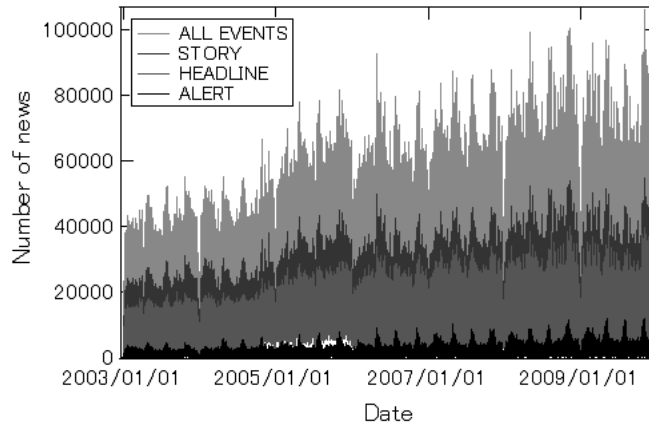


Fig. 1. Time series of the daily number of news reports. The number of news reported in English is counted. The top line is for all news; the second line is for story news; the third line is for headline news; and the bottom line is for alert news.

the frequency of news. Section 3 analyzes the autocorrelations of the frequency of news on a particular company and shows that the autocorrelation function follows a power law. Section 4 turns to the cross correlations for the frequency of news across companies. We show that the coupling of the average number of news on a company and its fluctuations obeys a scaling law, and that the frequency of news on a company is not governed solely by internal dynamics but is also affected by external dynamics, such as an increase in the number of news due to the outbreak of an economic crisis. In Section 5, we extract common movements across companies using random matrix theory techniques. Section 6 concludes the paper.

§2. Overview of the news data

Thomson Reuters Corporation is a world-famous provider of information for the world’s businesses and professionals, providing, among other things, “Reuters 3000 Xtra,” an electronic trading platform typically used by professional traders and investment analysts in trading rooms. “Reuters 3000 Xtra” offers real-time streaming news, comprehensive economic indicators, and financial data. It displays news of not only Thomson Reuters but also around 500 third parties. From 2003 to 2009, approximately 165 million news reports were provided. While these reports were in several languages, about 65 percent of them (107 million) were in English. In this study, we use only the English news reports, all of which are available in the Reuters News Scope Archive database.

There are three news types in the database. The first type is “alert” news, which cover an urgent newsworthy event and are 80-100 characters long. Alert news are normally followed by another news type. The second type is “headline” news, consisting of the headline of a news report for an event. The third type, finally, is “story” news, which bring the text that provides further information about the event. If the event is important, story news are often updated.

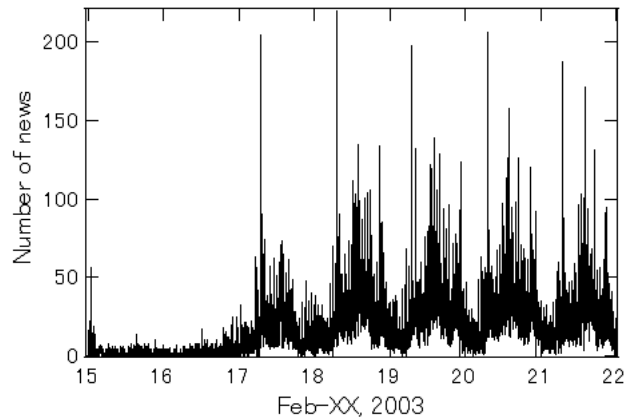


Fig. 2. Number of news reports per minute for a particular week, Feb. 15-22, 2003. The number of news reported in English is counted.

Table I. Mean of the number of news reports on each day of the week. The number of news reported in English is counted.

	Sat.	Sun.	Mon.	Tue.	Wed.	Thu.	Fri.
ALERT	101	148	3010	3807	3960	4366	2666
HEADLINE	2398	3967	20565	23115	23111	23453	20553
STORY	2614	4316	27811	31607	31633	32428	27804

Table II. Example of news

Date	Time	News type	Text
2010-04-05	00:03:14.307	STORY	...It topped Credit Suisse Group, which jumped from ninth place a year ago to second,...

Figure 1 shows the time series of the number of news reports for each news type. We find that the number of news delivered by Reuters increases every year. There were 9.8 million news reports in 2003, but 18.6 million in 2009. Figure 2 shows the time series of the number of news reports per minute for a particular week, namely that starting February 15, 2003. We clearly see that the frequency of news has intraday seasonality, as has also been observed in other studies.^{11),12)} Table 1 presents the mean of the number of news reports for each news type on each day of the week. There are fewer news reports on the weekend. This indicates that before proceeding to the detailed analysis we need to deal with the nonstationarity of the time series. How we do this is discussed in the next section.

In order to investigate the frequency of company news, we first need to construct time series for the number of news reports for each company. We do so using the following steps. First, we focus on the top 100 companies in the world in terms of market capitalization in 2003 and search the database by company name. For example, we find Credit Suisse Group mentioned in the text of a news report published at 00:03:14 on April 5 (see Table 2). Next, we define company news as news that mention the name of the company. Finally, by counting the number of news reports

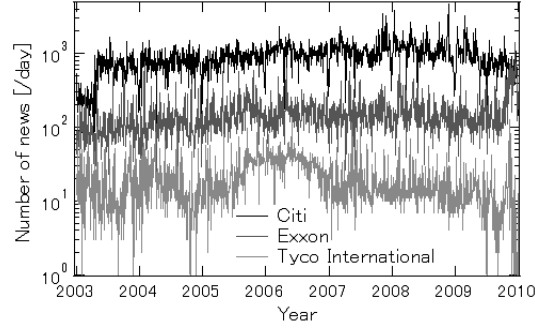


Fig. 3. Time series for the number of company news reports per day excluding reports on the weekend. The top, middle, and bottom lines are for news on Citi, Exxon, and Tyco International, respectively.

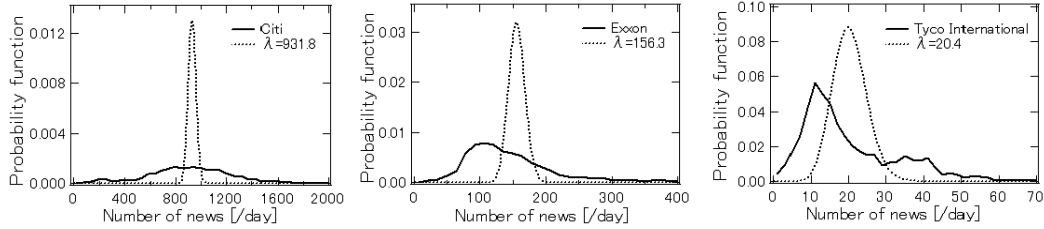


Fig. 4. Probability functions for the frequency of news on Citi, Exxon, and Tyco International. The dashed lines show Poisson distributions with the same mean as in the data. News reports on the weekend are excluded.

for each company, we obtain the time series.

§3. Autocorrelations for the frequency of news reports

We investigate the probability functions and autocorrelations of the frequency of company news for every news type. We focus on the time series of the number of news reports for three companies, Citi, Exxon, and Tyco International, which are shown in Figure 3. The daily mean of the number of reports on Citi, Exxon, and Tyco International, excluding reports on the weekend, is 932, 156, and 20, respectively. Figure 4 shows the probability function of the daily number of news reports on each company. Compared to a Poisson distribution with the same mean as in the data, each probability function has a fatter tail, suggesting that time series for company news do not follow a Poisson process.

The news frequency time series are not stationary due to the time trend and daily periodicity, as seen in the previous section. It may be the case that the fat tails of the probability functions observed in Figure 4 come from the nonstationarity of the time series. In order to transform our time series into stationary ones, we introduce the concept of “tick time” for news. Tick time refers not to actual time but is measured in terms of the appearance of news reports, where each new report corresponds to a unit of “time.” That is, tick time increases by one whenever a fresh news item

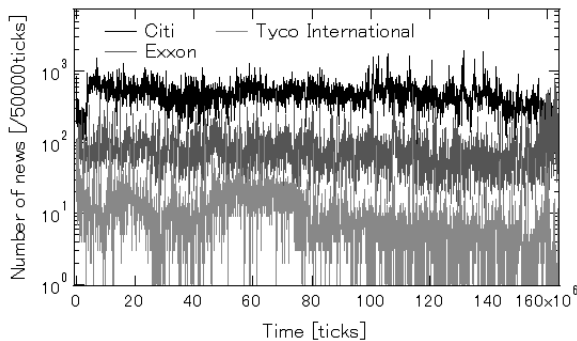


Fig. 5. Time series of the number of company news reports per 50,000 ticks. The top, middle, bottom lines are for Citi, Exxon, and Tyco International, respectively.

Table III. Augmented Dickey-Fuller test for the Citi, Exxon, and Tyco International series

		Trend	Drift	None
Citi	t -values	-19.7	-18.6	-1.08
	p -values	0.00	0.00	0.28
Exxon	t -values	-27.5	-26.8	-2.5
	p -values	0.00	0.00	0.01
Tyco International	t -values	-18.4	-18.0	-9.0
	p -values	0.00	0.00	0.00

in any language is reported. Note that because news reports are less frequent, and the interval between “ticks” in actual time therefore longer, on weekends, tick time passes more slowly on Saturdays and Sundays than during the daytime on weekends, when the number of news reports is large. We set tick time to zero for the beginning of our sample period (i.e., January 1st, 2003). Thus, using tick time allows us to eliminate the periodicity and trend observed in the original data. Figure 5 shows the time series of the number of news reports measured by tick time. In this figure, we count the number of news reports per 50,000 ticks, corresponding to about half a day, for each of the three companies. Comparing this with Figure 3, we see that the upward trend and daily periodicity have been eliminated.

To check the stationarity of the time series with tick time, we employ the Augmented Dickey-Fuller (ADF) test, which is a test for a unit root in a time series.^{13),14)} We choose the lag order of the ADF test using Akaike’s Information Criterion and conduct three types of ADF tests (“none,” “drift,” and “trend”) for the time series for Citi, Exxon, and Tyco International. If the type is set to “none,” neither an intercept nor a trend is included in the test regression; if it is set to “drift,” an intercept is added; and if it is set to “trend,” both an intercept and a trend are added. Table 3 presents the results of the ADF test for each type, showing that the null hypothesis that the time series measured by tick time is not stationary is rejected for eight cases out of the nine. In the rest of the paper, we will use tick time unless otherwise indicated.

We now turn to the estimation of the autocorrelation for the news frequency $f_{i,t}$

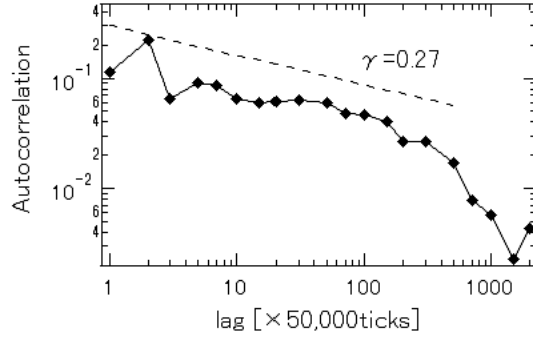


Fig. 6. Autocorrelation function of news reports. The frequency of news reports is defined as the number of news reports per 50,000 ticks, which is about half a day. The dashed reference line represents $\gamma = 0.27$.

of company i using an autocorrelation function of the form

$$\rho_i(\tau) = \frac{\langle f_{i,t} f_{i,t+\tau} \rangle - \langle f_{i,t} \rangle \langle f_{i,t+\tau} \rangle}{\sigma(f_{i,t}) \sigma(f_{i,t+\tau})} \quad (3.1)$$

where τ is a time lag, $\langle \cdot \rangle$ denotes the time average over the sample period, and σ is the standard deviation. We continue to measure the frequency by the number of news reports per 50,000 ticks. We pool observations for the top 100 companies. Figure 6 presents the estimated autocorrelation function, showing that it follows a power law of the form

$$\rho(\tau) \propto \tau^{-\gamma}. \quad (3.2)$$

where $\rho(\cdot)$ is the average of $\rho_i(\cdot)$ over the 100 companies. Note that the exponent γ is about 0.27, as represented by the reference line in the figure, and the estimated autocorrelation decays along the reference line up to $\tau = 600$, which is equivalent to about one year. This indicates the presence of long memory in the frequency of news reports. Such long memory properties were also observed for price volatility and trading volumes in stock markets (e.g., Ref. 15)).

§4. Scaling laws for the frequency of news

In this section, we investigate correlations in the frequency of news across different companies. A useful method to examine such cross correlations in the context of complex networks, such as the internet, is to look at the average flux and fluctuations on individual nodes.^{16)–18)} It was found that the coupling of the flux fluctuations with the total flux on individual nodes obeys a unique scaling law for a wide variety of complex networks, including the internet (i.e., a network of routers linked by physical connections), highways, river networks, and the World Wide Web of web pages and links.¹⁷⁾ Specifically, it was shown that the average flux $\langle f \rangle$ and the dispersion σ of those individual nodes are related as follows:¹⁷⁾

$$\sigma \sim \langle f \rangle^\alpha \quad (4.1)$$

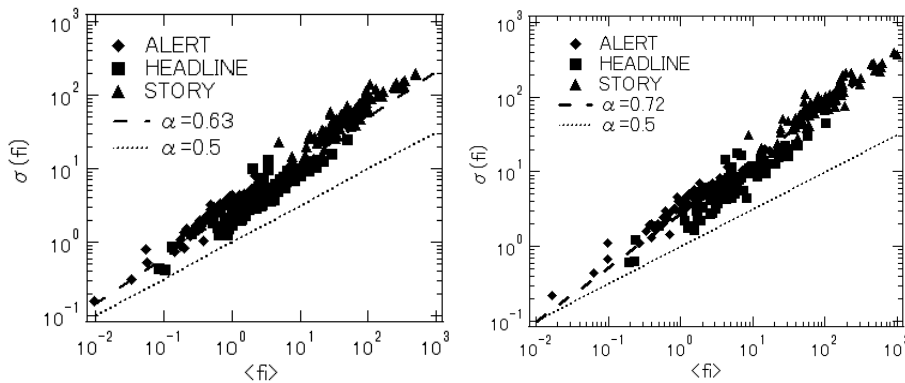


Fig. 7. The relationship between the mean and the standard deviation of the frequency of news reports for the top 100 companies in the world in terms of market capitalization in 2003. The frequency of news is defined as the number of news reports per 50,000 ticks in the left panel, while it is defined as the number of news per a day in the right panel.

where α is a scaling exponent. The scaling exponent α is equal to $1/2$ if the flux on individual nodes follows a Poisson process or is governed mainly by internal dynamics. On the other hand, the scaling exponent α deviates from $1/2$ if the flux does not follow a Poisson process and is equal to 1 if the flux on individual nodes is governed completely by external dynamics. For example, for river networks, the exponent α has been found to be quite close to unity, because the stream of rivers in different locations is mainly driven by weather patterns.

We apply this method to the frequency of news on individual companies by calculating the mean and standard deviation of the frequency of news for each company. Figure 7 plots $\sigma(f_i)$ for each of the top 100 companies in function of the average $\langle f_i \rangle$ of the same company. The frequency of news is defined as the number of news reports per 50,000 ticks in the left panel and as the number of news per day in the right panel. We see that in both cases the dots are not on the dotted line denoted by “ $\alpha = 1/2$ ”: the estimate of α is 0.63 in the case of tick time (left panel) and even higher in the case of actual time (right panel). These results suggest that the frequency of news is governed, at least partially, by external dynamics, such as the outbreak of an economic crisis that results in a simultaneous increase in the number of news for each company. Note that the higher estimate of α in the right panel can be interpreted as reflecting a closer comovement across companies due to intraday seasonality.

To see whether the scaling exponent α depends on the time scale, we estimate α for different time scales. Specifically, we count the number of news per s ticks, with s ranging from 5 to 100,000 ticks. Figure 8 shows that α is close to $1/2$ for small values of s , indicating that the frequency of news is governed by internal dynamics on small time scales such as minutes. However, α monotonically increases with time scale s and exceeds 0.6 for sufficiently large values of s , indicating that the frequency of news is governed, at least partially, by external dynamics on a daily or longer scale. Interestingly, a similar statistical property was found for transaction values on

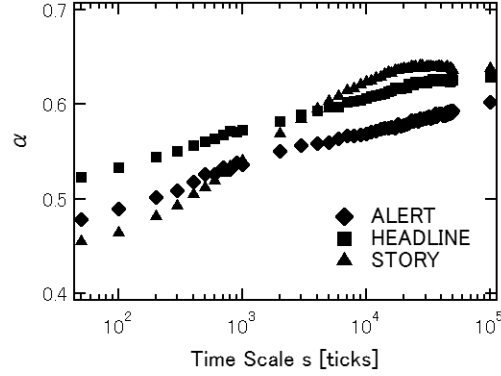


Fig. 8. Scaling exponent α for different time scales. The value of α is estimated using the observations for the top 100 companies in the world in terms of market capitalization in 2003.

the New York Stock Exchange; namely, α is close to $1/2$ on a time scale of minutes, while it is higher and close to unity on a daily or longer time scale.¹⁶⁾

§5. Extraction of common movements across companies

To learn more about the cross correlation detected in the previous section, we extract common movements across companies by applying random matrix theory (RMT) techniques to the cross-correlation matrix for the frequency of news reports. The cross-correlation matrix \mathbf{C} is defined by

$$C_{i,j} = \frac{\langle f_{i,t} f_{j,t} \rangle - \langle f_{i,t} \rangle \langle f_{j,t} \rangle}{\sigma(f_{i,t}) \sigma(f_{j,t})} \quad (5.1)$$

and is decomposed as

$$\mathbf{C} = \sum_{n=1}^N \lambda_n \mathbf{A}_n \mathbf{A}_n^T \quad (5.2)$$

where λ_n is the n -th largest eigenvalue and \mathbf{A}_n is the eigenvector associated with it. It has been shown that, if a cross-correlation matrix is generated from finite uncorrelated time series, the eigenvalue distribution of \mathbf{C} is given by

$$p(\lambda) = \begin{cases} \frac{Q}{2\pi} \frac{\sqrt{(\lambda_{\max} - \lambda)(\lambda - \lambda_{\min})}}{\lambda} & \text{if } \lambda_{\min} \leq \lambda \leq \lambda_{\max} \\ 0 & \text{otherwise} \end{cases} \quad (5.3)$$

where Q is defined as the ratio between the length of a time series L and the cross sectional dimension N (namely, $Q = L/N$), $\lambda_{\min} = \left(1 - \sqrt{1/Q}\right)^2$, and $\lambda_{\max} = \left(1 + \sqrt{1/Q}\right)^2$.^{19), 20)}

The sample period we analyze is seven years (January 2003 to December 2009), so that the length of time series L is 3,274 (i.e. $3,274 \times 50,000$ ticks). As before, we

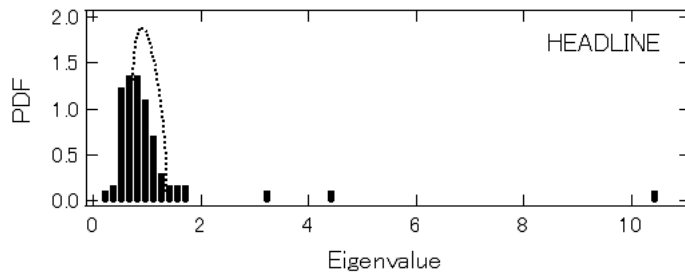


Fig. 9. The eigenvalue distribution for the case of headline news. This is estimated using the observations for the top 100 companies in the world in terms of market capitalization in 2003. The frequency of news is defined by the number of news reports per 50,000 ticks. The dotted line represents the eigenvalue distribution predicted in the case of finite uncorrelated time series, which is given by Eq. (5.3).

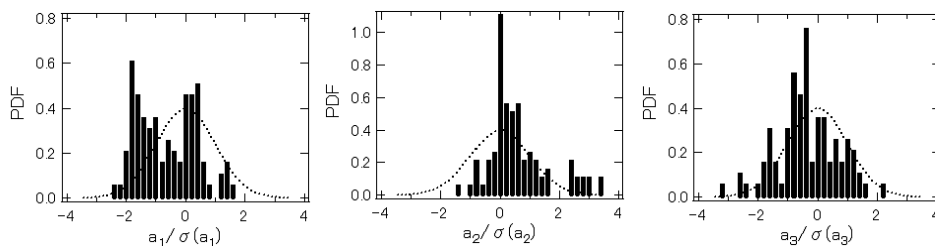


Fig. 10. The probability density functions of the eigenvector components associated with the first, second, and third largest eigenvalues. The horizontal axis shows the normalized component size (i.e., the size of components divided by the standard deviation). The dotted line represents the standard normal distribution, which is predicted in the case of finite uncorrelated time series.

pick the top 100 companies in terms of market capitalization in 2003. Given that $L = 3,274$ and $N = 100$, we have $\lambda_{\min} = 0.68$ and $\lambda_{\max} = 1.38$. Figure 9 shows the probability density function for the eigenvalues estimated from the cross-correlation matrix for the frequency of headline news, with the dotted line representing the eigenvalue distribution predicted in the case of finite uncorrelated time series, which is given by Eq. (5.3). There are eight eigenvalues exceeding λ_{\max} , with three of them exceeding λ_{\max} by a large margin.

Figure 10 presents the probability density functions for the eigenvector components associated with the largest, second largest, and third largest eigenvalues. We see that they deviate significantly from a standard normal distribution, which is predicted in the case of finite uncorrelated time series. Figure 11 shows the degree to which each company contributes to each of the eigenvectors associated with the three largest eigenvalues. The horizontal axis represents the 100 companies, which are sorted by industry codes. The three panels, each of which corresponds to the three largest eigenvalues, show that companies belonging to the financial sector greatly contribute to the eigenvector for the second largest eigenvalue (see the middle panel), and companies belonging to the information technology sector greatly contribute to the eigenvector for the third largest eigenvalue (the bottom panel). On the other hand, the top panel shows that almost all companies contribute evenly to

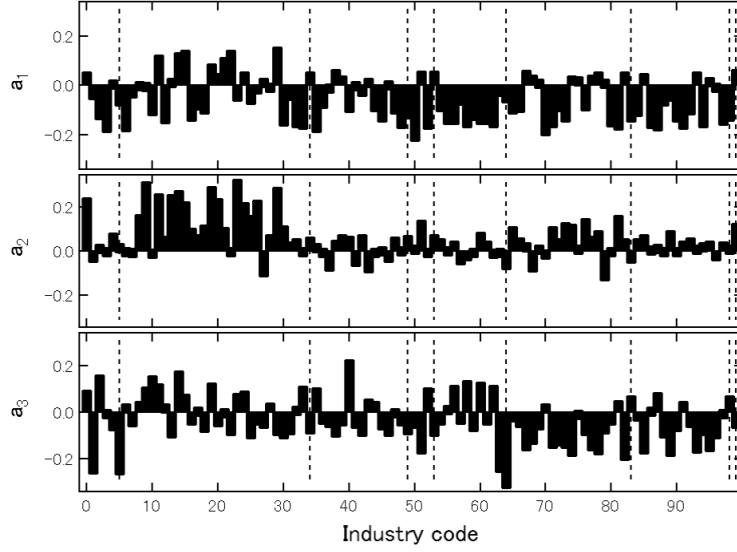


Fig. 11. Contributions of each company to the eigenvector components associated with the three largest eigenvalues of the correlation matrix. The upper, middle, and lower panels present the eigenvector components for the first, second, and third largest eigenvalues. The horizontal axis represents the 100 companies, which are sorted by industry code. Industry codes 0-4 represent basic materials industries; 5-33 financial services industries; 34-48 consumer goods industries; 49-52 conglomerates; 53-63 services industries; 64-82 information technology industries; 83-97 healthcare industries; 98 the industrial goods industry; and 99 utilities. The industry coding we use is available at http://biz.yahoo.com/ic/ind_index.html.

the eigenvector for the largest eigenvalue, which is similar to the result in Ref. 21) that the largest eigenvalue of the stock return correlation matrix is attributed to the “market mode” in financial markets.

Finally, we examine how the scaling exponent α changes when we eliminate a common movement across companies. We start by defining F_t as follows:

$$F_t = \sum_i a_{1,i} f_{i,t} \quad (5.4)$$

where $a_{1,i}$ denotes eigenvector component i for the largest eigenvalue. A similar variable was used to summarize a common movement for stock prices (see, e.g. Ref. 20)). We then eliminate the common movement by regressing $f_{i,t}$ on F_t :

$$f_{i,t} = b_i + d_i F_t + \epsilon_{i,t} \quad (5.5)$$

where b_i and d_i are regression coefficients. Using the residual term $\epsilon_{i,t}$ rather than $f_{i,t}$ itself, we estimate a scaling exponent α' satisfying a relationship of the form

$$\sigma(\epsilon_{i,t}) \propto \langle f_i \rangle^{\alpha'} \quad (5.6)$$

where $\sigma(\epsilon_{i,t})$ is the standard deviation of the residual term $\epsilon_{i,t}$. We find that the scaling exponent, which is equal to 0.63 when estimated using the original data, declines

to 0.61 when estimated after subtracting the common movement represented by F_t . This result suggests that the deviation of α from 1/2 shown in the previous section stems, at least partially, from the common movement across companies captured by the largest eigenvalue of the correlation matrix. It is natural to guess that the scaling exponent would approach to 1/2 when one further eliminates the common movements represented by the second largest eigenvalue, the third largest eigenvalue, and so on. It is our future task to see whether this is true or not by developing a method to eliminate the common movements represented by these eigenvalues.

§6. Conclusion

We empirically investigated temporal and cross correlations in the frequency of news reports on companies using a unique dataset with more than 100 million news articles reported in English by around 500 press agencies worldwide for the period 2003-2009. Our main findings are as follows. First, the frequency of news reports on a company does not follow a Poisson process; instead, it is characterized by long memory with a positive autocorrelation for more than a year. Second, there exist significant correlations in the frequency of news across companies. Specifically, on a daily or longer time scale, the frequency of news is governed by external dynamics such as an increase in the number of news due to, for example, the outbreak of an economic crisis, while it is governed by internal dynamics on a time scale of minutes. These two findings indicate that the frequency of news on companies has similar statistical properties as trading activities, measured by trading volumes or price volatility, in stock markets, suggesting that the flow of information through news on companies plays an important role in price dynamics in stock markets.

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